

Rich Dad, Smart Dad:

Decomposing the Intergenerational Transmission of Income

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Research Question

- Why is income correlated between fathers and sons?
- In particular...
- How much of the intergenerational income elasticity (IIE) can be attributed to the *causal* impact of fathers' income and how much can be attributed to the transmission of his human capital?

How do we study these questions?

- 2-factor model consistent with Becker & Tomes (1979)
 1. Fathers invest \$ in child quality
 - child quality \approx child's \$
 2. Fathers transfer human capital to their children
- We propose an IV method for bounding the structural parameters of this model

Estimating the IIE

- Equation (1)

$$inc_{son} = \beta_0 + \beta_1 inc_{father} + \varepsilon_{son}$$

- Estimates of the IIE for the U.S. \approx 0.4-0.6 (Solon 1992, Mazumder 2005)
- Estimates of the IIE for Sweden \approx 0.25 (Björklund et al. 1997, 2003, forthcoming)

The 2-factor model

- Equation (2)

$$inc_{father} = \gamma + HC_{father} + \eta_{father}$$

- Equation (3)

$$inc_{son} = \pi_0 + \pi_1 inc_{father} + \pi_2 HC_{father} + v_{son}$$

- Substituting (2) \rightarrow (3) yields Equation (4)

$$inc_{son} = \pi_0 + \pi_1\gamma + (\pi_1 + \pi_2)HC_{father} + \pi_1\eta_{father} + v_{son}$$

The empirical approach

- Given this model, the OLS slope estimator for equation (1), converges to:

$$p \lim \left(\hat{\beta}_1^{OLS} \right) = \pi_1 + \pi_2 \frac{\text{var} \left(HC_{father} \right)}{\text{var} \left(HC_{father} \right) + \text{var} \left(\eta_{father} \right)}$$

- Suppose that there exists a correlate of paternal income Z_{father}

$$p \lim \left(\hat{\beta}_1^{IV} \right) = \pi_1 + \pi_2 \frac{\text{cov} \left(HC_{father}, Z_{father} \right)}{\text{cov} \left(HC_{father}, Z_{father} \right) + \text{cov} \left(\eta_{father}, Z_{father} \right)}$$

- Different correlates identify a potentially different weighted combination of the structural parameters

The empirical approach: estimating bounds for π_1 and π_2 (method #1)

- Imagine a set of correlates of paternal income (instruments) that are correlated with luck and with human capital to varying degrees and satisfy the monotonicity condition that $\text{cov}(\text{HC}, Z)$ and $\text{cov}(\eta, Z)$ have the same sign
- The min IV estimate \rightarrow upper bound for π_1
- The max IV estimate \rightarrow lower bound for $\pi_1 + \pi_2$
- Max – min \rightarrow lower bound for π_2

The empirical approach: estimating bounds for π_1 and π_2 (method #2)

- Imagine a set of variables that are correlated with human capital and **not** correlated with luck
- IV estimate identifies $\pi_1 + \pi_2$
- Recall that

$$p \lim \left(\hat{\beta}_1^{OLS} \right) = \pi_1 + \pi_2 \frac{\text{var} \left(HC_{father} \right)}{\text{var} \left(HC_{father} \right) + \text{var} \left(\eta_{father} \right)}$$

method #2...contd.

- With an estimate of $\text{var}(\text{HC})/[\text{var}(\text{HC}) + \text{var}(\eta)]$ we can recover π_1 and π_2
- Mincer R^2 yields a lower bound for this variance share
- With mincer R^2 + OLS estimate + IV estimate in hand we can recover a lower bound for π_2 and an upper bound for π_1

Data sources

- Sweden's multigenerational register → identifies fathers and their sons
- Tax register → pre-tax total factor income 1968 – 2005
- Census data; 1960, 65, 70, 75, 80, 85, 90
 - Employment status
 - Municipality of residence
 - Occupation
 - Education (for some older fathers)
- Education register

Measuring permanent income

- Fathers
 - Log average income age 30 – age 60
 - At least 10 non-missing observations
- Sons
 - Log average income age 30 – age 40
 - At least 10 non-missing observations

Our Instruments

- Human capital
 - Education; years of schooling or 7 categories
 - Occupation; initial occupation in 1970 or occupation 1970 - 1990
- Luck
 - Employment status; dichotomous variable indicating good or bad state
 - Employment residuals; employment status \perp education and past income

Baseline estimate of father-son IIE

- Log average income
- IIE = 0.286 (0.010)

IV estimates

- Human capital
 - Years of education = 0.416 (0.020)
 - Education category = 0.414 (0.020)
 - Occupation 1970 = 0.400 (0.014)
 - Occupation 1970-90 = 0.335 (0.011)
- Luck
 - Father employment status = 0.205 (0.017)
 - Father employment residuals = 0.106 (0.065)

Decomposition: method #1

- OLS IIE = 0.286
- $\pi_1 = 0.106$
 - from employment residuals IV
- $\pi_1 + \pi_2 = 0.414$
 - education category IV
- $\pi_2 = 0.307$
 - implied by model

Decomposition: method #1

- At most 37% of the IIE reflects the *causal* effect of fathers' financial resources
 - Upper bound
- The remainder captures the impact of fathers' human capital

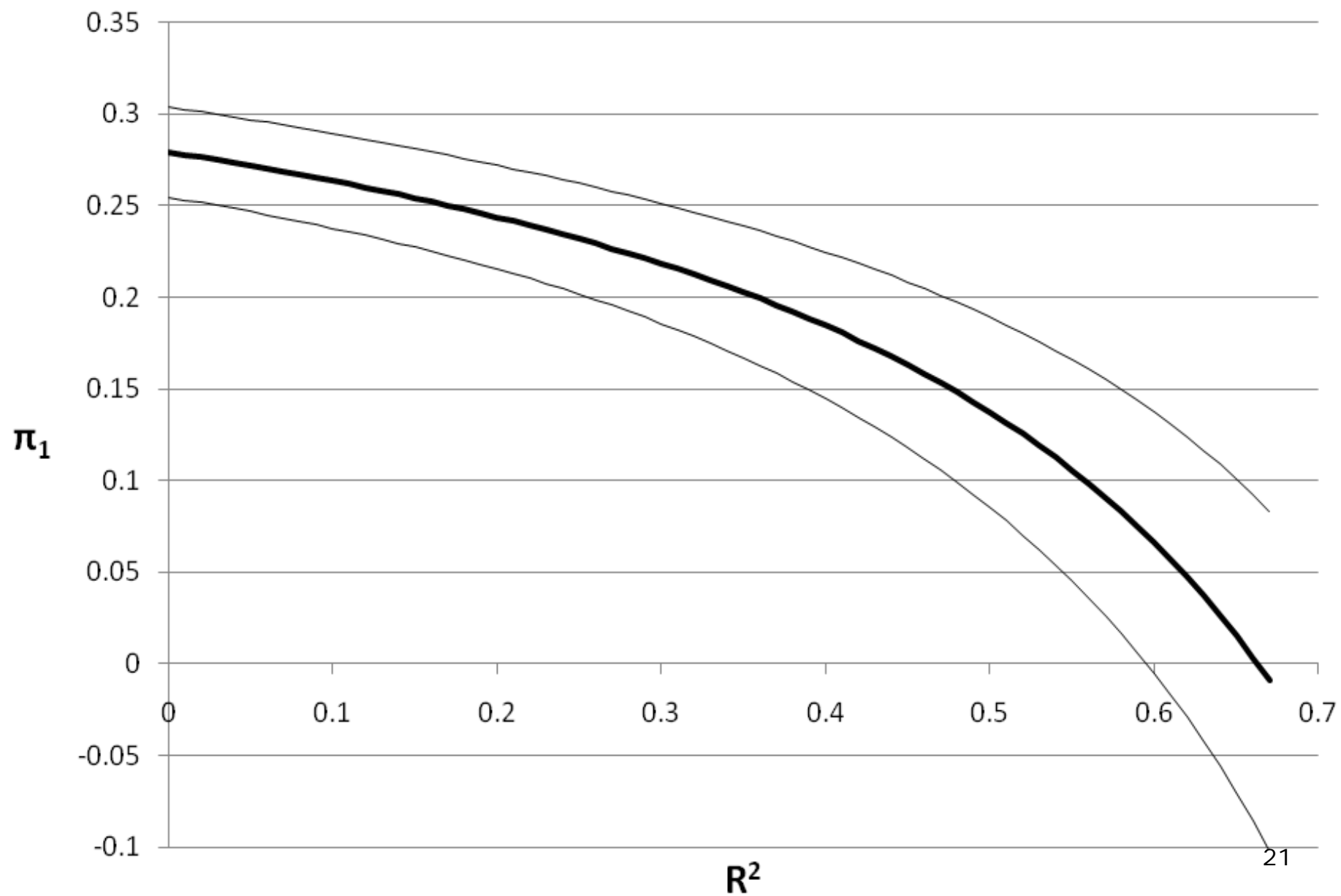
Decomposition: method #2

- OLS IIE = 0.286
- $\pi_1 + \pi_2 = 0.414$
 - education category IV
- Mincer $R^2 = 0.376$
 - estimated from sample
- $\pi_1 = 0.209$
 - implied by model
- $\pi_2 = 0.205$
 - implied by model

Decomposition: method #2

- At most 72% of the IIE reflects the *causal* effect of fathers' financial resources
 - Upper bound
- The remainder captures the impact of fathers' human capital

π_1 as a Function of Mincer R^2



Robustness checks

- Alternative income measures
- Spatial correlations + regional price indices
- Non-linearities
- 3-factor model

Conclusion

- Summary of results
 1. Swedish IIE = 0.286
 2. At most, 37% is due to the *causal* impact of fathers' financial resources
 3. The remainder is due to the transmission of fathers' human capital
- Possible extensions
 1. Cross-country comparisons
 2. Extend to a 3-factor model; \$, genetic HC, non-genetic HC
 3. ...with and without assortative mating